

A Framework for Bounding Deterministic Risk with PAC-Bayes: Applications to Majority Votes

Benjamin Leblanc

<https://benjaminlblanc.github.io/>

Université Laval, département d'informatique et de génie logiciel

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1 Preamble

- Definitions
- Generalization guarantees / PAC-Bayes

2 Derandomizing PAC-Bayes bounds - Majority votes

- Categorical distribution
- Dirichlet distribution
- Gaussian distribution
- Empirical results

Definitions

An **observation** $\mathbf{z} := (\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y} = \mathcal{Z}$ is a pair **features** - **labels**.

Type of task

$$\mathcal{X} \subseteq \mathbb{R}^d, \mathcal{Y} \in \{\mathbf{y} \in \{0,1\}^c \mid \sum_{i=1}^c y_i = 1\}$$

Data generating distribution

Examples comes from a **distribution** D over $\mathcal{Z}$.

Training set

$$S := \{\mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_m\} \sim D^m$$

Classifier (hypothesis), hypothesis space

$$h : \mathcal{X} \rightarrow \mathcal{Y}', \quad \mathcal{Y}' \supseteq \mathcal{Y}$$

Hypothesis space

$$h \in \mathcal{H}$$

Loss function

$$\ell : \mathcal{Y}' \times \mathcal{Y} \rightarrow \{0,1\}$$

Definitions

Training loss

$$\widehat{\mathcal{L}}_S(h) := \frac{1}{n} \sum_{i=1}^m \ell(h(\mathbf{x}_i), y_i)$$

Goal : minimize the average loss on unseen examples (generalization loss, risk)

$$\mathcal{L}_D(h) := \mathbb{E}_{z \sim D} \ell(h(\mathbf{x}), y)$$

→ Throughout the presentation, for any quantity K_D that depends on D :

- $\widehat{K}_S$ represents the empirical estimation of K_D from S ;
- $\widetilde{K}_S$ represents a bound on K_D using S .

The generalization problem

PAC (Probably Approximately Correct) bounds

"With probability at least $1-\delta$ ", the generalization loss of h will be at most $\epsilon(\cdot, \dots)$ "

$$\mathbb{P}_{S \sim D^m} \left(\mathcal{L}_D(h) \leq \epsilon(\hat{\mathcal{L}}_S(h), m, \delta, \dots) \right) \geq 1-\delta$$

Why PAC-Bayes ?

PAC (Probably Approximately Correct)

With probability at least “ $1-\delta$ ”, the generalization loss of h will be at most “ $\varepsilon(\cdot, \dots, \cdot)$ ”

$$\mathbb{P}_{S \sim D^m} \left(\mathcal{L}_D(h) \leq \varepsilon(\hat{\mathcal{L}}_S(h), m, \delta, \dots) \right) \geq 1-\delta$$

- A single hypothesis ($\mathcal{H} = \{h\}$) :

$$\mathcal{L}_D(h) \leq \hat{\mathcal{L}}_S(h) + \sqrt{\frac{1}{2m} \log \left(\frac{1}{\delta} \right)}.$$

- Finite hypothesis space ($|\mathcal{H}| < \infty$) :

$$\forall h \in \mathcal{H}, \quad \mathcal{L}_D(h) \leq \hat{\mathcal{L}}_S(h) + \sqrt{\frac{1}{2m} \log \left(\frac{|\mathcal{H}|}{\delta} \right)}$$

- Countable hypothesis space with probability *a priori* $p(h_i)$:

$$\forall h \in \mathcal{H}, \quad \mathcal{L}_D(h) \leq \hat{\mathcal{L}}_S(h) + \sqrt{\frac{1}{2m} \log \left(\frac{1}{p(h)\delta} \right)}$$

- Uncountable hypothesis space : VC dimension, Rademacher complexity, PAC-Bayes...

Typical PAC-Bayes

PAC-Bayesian considerations

We will define our bounds on **average** over the random draw of a hypothesis following an **a posteriori distribution** Q over $\mathcal{H}$.

$$\mathbb{P}_{S \sim D^m} \left(\underbrace{\mathbb{E}_{h' \sim Q} \mathcal{L}_D(h')}_{:= \mathcal{L}_D(Q), \text{ "Gibbs risk" }} \leq \dots \right) \geq 1 - \delta$$

→ Throughout the presentation :

- h' is always drawn from Q , whereas h is chosen.

Major con of such an approach ?

We do not have any information on any single predictor ! Not usable in many applications...

1 Preamble

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2 Derandomizing PAC-Bayes bounds - Majority votes

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Derandomizing PAC-Bayes Bounds

The question of interest : Is it possible to bound a *single, chosen* classifier of an *uncountable hypothesis space* using a PAC-Bayes bound ?

$$\text{i.e. : } \mathcal{L}_D(h) \leq f(\mathcal{L}_D(Q)) \leq f(\widetilde{\mathcal{L}}_D(Q)) \quad \mathcal{L}_D(h) \leq \underbrace{f(\mathcal{L}_D(Q)) \leq f(\widetilde{\mathcal{L}}_D(Q))}_{\text{This is what PAC-Bayes does}} \quad \overbrace{\mathcal{L}_D(h) \leq f(\mathcal{L}_D(Q))}^{\text{This is what we want !}} \leq \underbrace{f(\mathcal{L}_D(Q))}_{\text{This is what we want !}}$$

Proposition 1.

For any data distribution D , distribution Q over $\mathcal{H}$, classifier $h \in \mathcal{H}$: if $c_Q(h) \neq b_Q(h)$, then

$$\mathcal{L}_D(h) = \mathcal{L}_D(h) = \frac{\mathcal{L}_D(Q) \mathcal{L}_D(Q) - b_Q(h)}{c_Q(h) - b_Q(h)},$$

where

$$c_Q(h) \underbrace{c_Q(h)} = [\dots], \quad b_Q(h) \underbrace{b_Q(h)} = [\dots].$$

Involve both Q and h .

A first step

Proposition 1. *proof sketch.*

$$\begin{aligned}\mathcal{L}_D(Q) &= \mathbb{E}_{(\mathbf{x}, y) \sim D} \mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \\&= \mathbb{P}_{(\mathbf{x}, y) \sim D} \left(\ell(h(\mathbf{x}), y) = 1 \right) \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 1 \right] + \\&\quad \mathbb{P}_{(\mathbf{x}, y) \sim D} \left(\ell(h(\mathbf{x}), y) = 0 \right) \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 0 \right] \\&= \underbrace{\mathbb{P}_{(\mathbf{x}, y) \sim D} \left(\ell(h(\mathbf{x}), y) = 1 \right)}_{\mathcal{L}_D(h)} \underbrace{\mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 1 \right]}_{c_Q(h)} + \\&\quad \underbrace{\mathbb{P}_{(\mathbf{x}, y) \sim D} \left(\ell(h(\mathbf{x}), y) = 0 \right)}_{1 - \mathcal{L}_D(h)} \underbrace{\mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 0 \right]}_{b_Q(h)}\end{aligned}$$

A first step

Proposition 1.

For any data distribution D , distribution Q over $\mathcal{H}$, classifier $h \in \mathcal{H}$: if $c_Q(h) \neq b_Q(h)$, then

$$\mathcal{L}_D(h) = \frac{\mathcal{L}_D(Q) - b_Q(h)}{c_Q(h) - b_Q(h)},$$

where

$$c_Q(h) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 1 \right], \quad b_Q(h) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 0 \right].$$

Corollary 1.

Let $\mathcal{L}_D(Q) \leq \widetilde{\mathcal{L}}_S(Q)$ ✓, $c_Q(h) \geq \widetilde{c}_S(h)$?, and $b_Q(h) \geq \widetilde{b}_S(h)$? : if $\widetilde{c}_S(h) > \widetilde{b}_S(h)$?, then

$$\mathcal{L}_D(h) \leq \frac{\widetilde{\mathcal{L}}_S(Q) - \widetilde{b}_S(h)}{\widetilde{c}_S(h) - \widetilde{b}_S(h)}.$$

A first step

Why not simply use regular PAC-Bayesian results to bound $c_Q(h)$ and $b_Q(h)$?

Lemma 1.

Let $\text{kl}(q, p) = q \ln(\frac{q}{p}) + (1 - q) \ln(\frac{1-q}{1-p})$. For any distribution D , hypothesis set $\mathcal{H}$, prior distribution P over $\mathcal{H}$, $\delta \in (0, 1]$, we have with probability at least $1 - \delta$ over the random choice $S \sim D^m$ that for every Q over $\mathcal{H}$:

$$\text{kl} \left(\widehat{\mathcal{L}}_S(Q) \parallel \mathcal{L}_D(Q) \right) \leq \frac{1}{m} \left[\text{KL}(Q \parallel P) + \ln \left(\frac{2\sqrt{m}}{\delta} \right) \right].$$

$$\text{kl} \left(\widehat{b}_S(h) \parallel b_Q(h) \right) \leq \frac{1}{m} \left[\text{KL}(Q \parallel P) + \ln \left(\frac{2\sqrt{m}}{\delta} \right) \right] ?$$

$$\rightarrow b_Q(h) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 0 \right] :$$

A first step

PAC-Bayes won't be of help here... We need the bound to hold simultaneously for all $h \in \mathcal{H}$!

We need a way to :

- Bound $b_Q(h) \geq \widetilde{b}_S(h)$ and $c_Q(h) \geq \widetilde{c}_S(h)$ with probability 1 ;
- Ensure that $\widetilde{c}_S(h) > \widetilde{b}_S(h)$.

Possible solution : bound $b_Q(h)$ and $c_Q(h)$ by their worst-case values ! In there current form

$\left(\text{i.e. } \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{h' \sim Q} \ell(h'(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 0.5 \pm 0.5 \right] \right)$, hard to grasp what their worst value could be ; missing link between h and Q ...

Solution : we focus on specific $\mathcal{H}$ and Q , so that simplification might occur !

A specialization to majority votes

Base classifiers

$$\mathcal{F} = \{f_1, \dots, f_n\}, f_i : \mathcal{X} \rightarrow \mathcal{Y} \forall i \in \{1, \dots, n\}$$

Hypothesis space

$$\mathcal{H} = \{\sum_{i=1}^n \alpha_i f_i \mid \alpha \in \mathcal{A}\}, \mathcal{A} \subseteq \mathbb{R}^n$$

Majority vote

$$h_\alpha(\mathbf{x}) = \sum_{i=1}^n \alpha_i f_i(\mathbf{x}) = \alpha \cdot \mathbf{f}(\mathbf{x}) \in \mathcal{Y}'$$

→ For simplicity, we now define Q over the weight space $\mathcal{A}$: each vector of weights α corresponds to a unique hypothesis h_α .

Gibbs risk (for any Q over $\mathcal{A}$)

$$\mathcal{L}_D(Q) := \mathbb{E}_{z \sim D} \mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y) = \mathbb{E}_{\alpha \sim Q} \mathcal{L}_D(h_\alpha)$$

Theory under the Categorical assumption

We first explore the most restrictive support for $Q : \{\alpha \in \{0, 1\}^n \mid \sum_{i=1}^n \alpha_i = 1\}$. A natural choice for Q is the Categorical distribution, and a natural choice for ℓ is

$$\ell(h_\alpha(\mathbf{x}), y) = \mathbb{1} \left\{ \sum_{i=1}^n \alpha_i \mathbb{1} \{f_i(\mathbf{x}) \neq y\} \underbrace{\sum_{i=1}^n \alpha_i \mathbb{1} \{f_i(\mathbf{x}) \neq y\}}_{\alpha(\mathbf{x}, y)} \geq 0.5 \right\}.$$

Theory under the Categorical assumption

Let us recall that $Q \sim \mathcal{C}(\mathbf{p})$. Given $(\mathbf{x}, y)$, we have

$$c_Q(h) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y) \mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y) \mid \ell(h(\mathbf{x}), y) = 1 \right].$$

We first focus on the inner content of $c_Q(h)$ and $b_Q(h)$: $\mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y)$.

Theory under the Categorical assumption

Remember that $Q \sim \mathcal{C}(\mathbf{pp})$.

$$\begin{aligned}\mathbb{E}_{\alpha \sim Q} \ell(h_{\alpha}(\mathbf{x}), y) &= \mathbb{E}_{\alpha \sim Q} \mathbb{1} \left\{ \sum_{j=1}^n \alpha_j \mathbb{1} \{f_j(\mathbf{x}) \neq y\} \geq 0.5 \right\} \\&= \sum_{i=1}^n p_i \mathbb{1} \{ \mathbb{1} \{f_i(\mathbf{x}) \neq y\} \geq 0.5 \} \\&= \sum_{i=1}^n p_i \mathbb{1} \{f_i(\mathbf{x}) \neq y\} \\&= \mathbf{pp}(\mathbf{x}, y)\end{aligned}$$

Theory under the Categorical assumption

$$\begin{aligned}c_Q(h.) &= \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[\mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y) \mid \ell(h.(\mathbf{x}), y) = 1 \right] \\&= \mathbb{E}_{(\mathbf{x}, y) \sim D} [\mathbf{p}(\mathbf{x}, y) \mid \ell(h.(\mathbf{x}), y) = 1] \\&= \mathbb{E}_{(\mathbf{x}, y) \sim D} [\mathbf{p}(\mathbf{x}, y) \mid \ell(h_{\mathbf{p}}(\mathbf{x}), y) = 1] && \langle \text{By choosing } h. := h_{\mathbf{p}} \rangle \\&= \mathbb{E}_{(\mathbf{x}, y) \sim D} [\mathbf{p}(\mathbf{x}, y) \mid \mathbf{p}(\mathbf{x}, y) \geq 0.5]\end{aligned}$$

Theory under the Categorical assumption

Proposition 1.1.

Let $Q \sim \mathcal{C}(\mathbf{p})$ be a Categorical distribution with parameters $\mathbf{p}$. For any data distribution D , $\mathbf{p} \in [0, 1]^n$ s.t. $\sum_{i=1}^n p_i = 1$:

$$c_Q(h_{\mathbf{p}}) = \mathbb{E}_{(\mathbf{x}, y) \sim D} [\mathbf{p}(\mathbf{x}, y) \mid \mathbf{p}(\mathbf{x}, y) \geq 0.5] , \quad b_Q(h_{\mathbf{p}}) = \mathbb{E}_{(\mathbf{x}, y) \sim D} [\mathbf{p}(\mathbf{x}, y) \mid \mathbf{p}(\mathbf{x}, y) < 0.5] ,$$

where $\mathbf{p}(\mathbf{x}, y) = \sum_{i=1}^n p_i \mathbb{1}\{f_i(\mathbf{x}) \neq y\}$.

We need a way to :

- Bound $b_Q(h) \geq \widetilde{b}_S(h)$ and $c_Q(h) \geq \widetilde{c}_S(h)$ with probability 1 ;
- Ensure that $\widetilde{c}_S(h) > \widetilde{b}_S(h)$. ✓

Theory under the Categorical assumption

We seek upper bounds on $b_Q(h_{\mathbf{p}})$ and $c_Q(h_{\mathbf{p}})$ that hold with probability 1.

Given $\mathbf{p}$:

$$\begin{aligned} b_Q(h_{\mathbf{p}}) &= \mathbb{E}_{(\mathbf{x}, y) \sim D} [p(\mathbf{x}, y) \mid p(\mathbf{x}, y) < 0.5] \\ &\geq \min_{(\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y}} [p(\mathbf{x}, y) \mid p(\mathbf{x}, y) < 0.5] \\ &= \min_{(\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y}} [p(\mathbf{x}, y)] \end{aligned}$$

→ Easy to compute for many hypotheses.

Theory under the Categorical assumption

Similarly :

$$\begin{aligned} c_Q(h_{\mathbf{p}}) &= \mathbb{E}_{(\mathbf{x}, y) \sim D} [p(\mathbf{x}, y) \mid p(\mathbf{x}, y) \geq 0.5] \\ &\geq \min_{(\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y}} [p(\mathbf{x}, y) \mid p(\mathbf{x}, y) \geq 0.5] \\ &\geq \min_{\beta \in \{0,1\}^n} [\beta \cdot \mathbf{p} \mid \beta \cdot \mathbf{p} \geq 0.5] . \end{aligned}$$

Let's make an example : suppose $\mathbf{p} = [0.33, 0.33, 0.33]$. What is $\min_{\beta \in \{0,1\}^n} [\beta \cdot \mathbf{p} \mid \beta \cdot \mathbf{p} \geq 0.5]$?

Now suppose $|\mathbf{p}| > 1000$! There are $O(2^n)$ possible values for $\beta \cdot \mathbf{p}$...

Knapsack problem representation non efficient : $\mathbf{p}$ is a real-valued vector...

Solution : partition problem representation !

Theory under the Categorical assumption

The partition problem

Given a set of number $\mathbf{p}$, the partition problem consists in the following :

$$\text{Part}(\mathbf{p}) = \underset{\mathbf{p}_1, \mathbf{p}_2}{\operatorname{argmin}} \{ |\text{sum}(\mathbf{p}_1) - \text{sum}(\mathbf{p}_2)| \text{ s.t. } \mathbf{p}_1 \cup \mathbf{p}_2 = \mathbf{p} \wedge \mathbf{p}_1 \cap \mathbf{p}_2 = \emptyset \}.$$

Proposition 1.1.1.

Given a set of number $\mathbf{p} \in [0, 1]^n$, where $\text{sum}(\mathbf{p}) = 1$, let $\text{Part}(\mathbf{p}) = \mathbf{p}_1, \mathbf{p}_2$. Then :

$$\min_{\beta \in \{0,1\}^n} [\beta \cdot \mathbf{p} \mid \beta \cdot \mathbf{p} \geq 0.5] = \max(\text{sum}(\mathbf{p}_1), \text{sum}(\mathbf{p}_2)).$$

Indeed : $\text{Part}([0.33, 0.33, 0.33]) = [0.33], [0.33, 0.33]$.

→ There exist complete greedy algorithms for the partitioning problem !

Theory under the Categorical assumption

Proposition 1.1.1.

Given a set of number $\mathbf{p} \in [0, 1]^n$, where $\text{sum}(\mathbf{p}) = 1$, let $\text{Part}(\mathbf{p}) = \mathbf{p}_1, \mathbf{p}_2$. Then :

$$\min_{\beta \in \{0,1\}^n} [\beta \cdot \mathbf{p} \mid \beta \cdot \mathbf{p} \geq 0.5] = \max(\text{sum}(\mathbf{p}_1), \text{sum}(\mathbf{p}_2)).$$

We need a way to :

- Bound $b_Q(h) \geq \widetilde{b}_S(h)$ and $c_Q(h) \geq \widetilde{c}_S(h)$ with probability 1 ; ✓
- Ensure that $\widetilde{c}_S(h) > \widetilde{b}_S(h)$. ✓

Corollary 1.

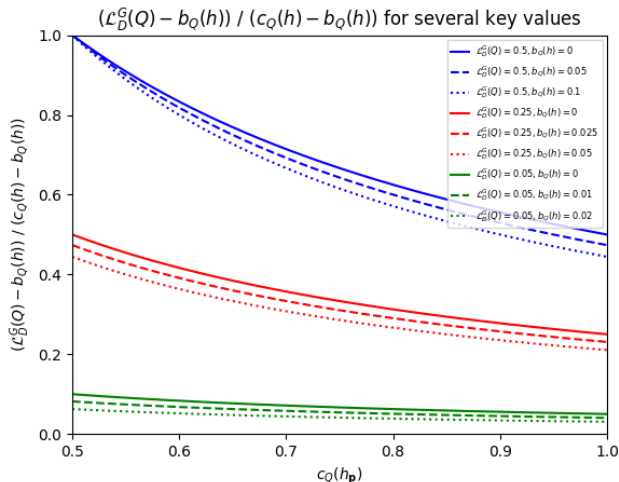
Let $\mathcal{L}_D(Q) \leq \widetilde{\mathcal{L}}_S(Q)$, $c_Q(h_{\mathbf{p}}) \geq \widetilde{c}_S(h_{\mathbf{p}})$, and $b_Q(h_{\mathbf{p}}) \geq \widetilde{b}_S(h_{\mathbf{p}})$: if $\widetilde{c}_S(h_{\mathbf{p}}) > \widetilde{b}_S(h_{\mathbf{p}})$, then

$$\mathcal{L}_D(h_{\mathbf{p}}) \leq \frac{\widetilde{\mathcal{L}}_S(Q) - \widetilde{b}_S(h_{\mathbf{p}})}{\widetilde{c}_S(h_{\mathbf{p}}) - \widetilde{b}_S(h_{\mathbf{p}})}.$$

Theory under the Categorical assumption

Worst-case scenario :

- $\widetilde{b}_S(h_p) = 0$,
- $\widetilde{c}_S(h_p) = 0.5$:
- $\frac{\mathcal{L}_D(Q) - \widetilde{b}_S(h_p)}{\widetilde{c}_S(h_p) - \widetilde{b}_S(h_p)} = 2\mathcal{L}_D(Q).$



Theory under the Dirichlet assumption

We now explore a lesser restrictive support for $Q : \{\alpha \in [0, 1]^n \mid \sum_{i=1}^n \alpha_i = 1\}$. A natural choice for Q is the Dirichlet distribution.

Proposition 1.2.

Let $\mathcal{Y}' = \mathcal{Y}$, $\ell(h_\alpha(\mathbf{x}), y) = \mathbb{1} \{ \sum_{i=1}^n \alpha_i \mathbb{1} \{ f_i(\mathbf{x}) \neq y \} \geq 0.5 \}$ and $Q \sim \mathcal{D}(\mathbf{p})$ be a Dirichlet distribution with parameters $\mathbf{p}$. For any data distribution D , $\mathbf{p} \in \mathbb{R}_{>0}^n$:

$$c_Q(h_{\mathbf{p}}) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[l_{0.5}(\bar{\mathbf{p}} - \mathbf{p}(\mathbf{x}, y), \mathbf{p}(\mathbf{x}, y)) \mid \frac{\mathbf{p}(\mathbf{x}, y)}{\bar{\mathbf{p}}} \geq 0.5 \right],$$
$$b_Q(h_{\mathbf{p}}) = \mathbb{E}_{(\mathbf{x}, y) \sim D} \left[l_{0.5}(\bar{\mathbf{p}} - \mathbf{p}(\mathbf{x}, y), \mathbf{p}(\mathbf{x}, y)) \mid \frac{\mathbf{p}(\mathbf{x}, y)}{\bar{\mathbf{p}}} < 0.5 \right],$$

where $l_x(\cdot, \cdot)$ is the regularized incomplete beta function evaluated at x , and $\bar{\mathbf{p}} = \sum_{i=1}^n a_i$.

$\rightarrow c_Q(h_{\mathbf{p}})$ and $b_Q(h_{\mathbf{p}})$ are proportional to $\mathbf{p}(\mathbf{x}, y)!$

Proposition 1.2. *proof sketch*

Based on ZANTEDESCHI et al. 2021, we use the following result :

$$\mathbb{E}_{\alpha \sim \mathcal{D}(\mathbf{p})} \mathbb{1} \left\{ \sum_{i=1}^n \alpha_i \mathbb{1} \{f_i(\mathbf{x}) \neq y\} \geq 0.5 \right\} = l_{0.5}(\bar{\mathbf{p}} - \mathbf{p}(\mathbf{x}, y), \mathbf{p}(\mathbf{x}, y)).$$

Theory under the Gaussian assumption

We now explore the most general support for $Q : \mathbb{R}^n$.

A natural choice for Q is the Gaussian distribution. What loss to use? Remember : we must have an analytic form for $\mathbb{E}_{\alpha \sim Q} \ell(h_\alpha(\mathbf{x}), y) \dots$

Such a result only exists for $\mathcal{Y} = \{-1, +1\}$ and $\ell(h_\alpha(\mathbf{x}), y) = \mathbb{1} \{y \sum_{i=1}^n \alpha_i f_i(\mathbf{x}) \leq 0\}$.

Proposition 1.3.1

Let $\mathcal{Y} = \{-1, +1\}$ and $\mathcal{Y}' \subseteq \mathbb{R}$. Let $\ell(h_\alpha(\mathbf{x}), y) = \mathbb{1} \{y \sum_{i=1}^n \alpha_i f_i(\mathbf{x}) \leq 0\}$ and $Q \sim \mathcal{N}(\boldsymbol{\mu}, I)$ be an isotropic Gaussian distribution with mean $\boldsymbol{\mu}$. For any data distribution D , $\boldsymbol{\mu} \in \mathbb{R}^n$:

$$c_Q(h_\mu) = \mathbb{E}_{(\mathbf{x}, \cdot) \sim D} \left[1 - \Phi \left(\frac{|\boldsymbol{\mu} \cdot \mathbf{f}(\mathbf{x})|}{\|\mathbf{f}(\mathbf{x})\|} \right) \right],$$

$$b_Q(h_\mu) = \mathbb{E}_{(\mathbf{x}, \cdot) \sim D} \left[\Phi \left(\frac{|\boldsymbol{\mu} \cdot \mathbf{f}(\mathbf{x})|}{\|\mathbf{f}(\mathbf{x})\|} \right) \right],$$

where $\Phi(k) = \frac{1}{2} \left(1 - \frac{2}{\sqrt{\pi}} \int_0^{k/\sqrt{2}} e^{-t^2} dt \right)$.

Empirical results

- Tested on 11 binary/multiclass datasets
- Using Seeger's bound to bound the Gibbs risk
- Base classifier are :
 - Uniformly spaced decision stumps (binary classification), or
 - Decision trees learned on half the examples (multiclass classification)
- Five benchmarks from the literature

Empirical results

Table – Comparison of our proposed approach and a few baselines for different types of predictors.
Bolded values correspond to the best bound values. Average and standard deviations computed over 5 random seeds.

Task	FO		SO		Bin		CBnd		VC-dim		Part	
	Bound	Test error	Bound	Test error	Bound	Test error	Bound	Test error	Bound	Test error	Bound	Test error
ADULT	46.9 $\pm$ 0.0	22.2 $\pm$ 0.1	52.9 $\pm$ 0.1	15.8 $\pm$ 0.2	51.4 $\pm$ 0.0	24.1 $\pm$ 0.0	74.5 $\pm$ 0.4	21.0 $\pm$ 1.6	78.9 $\pm$ 0.0	17.2 $\pm$ 0.3	23.2 $\pm$ 2.4	21.7 $\pm$ 2.5
CODRNA	40.2 $\pm$ 0.1	12.0 $\pm$ 0.1	47.4 $\pm$ 0.1	11.7 $\pm$ 0.1	42.8 $\pm$ 5.2	15.2 $\pm$ 4.5	63.7 $\pm$ 1.0	23.1 $\pm$ 1.3	38.8 $\pm$ 1.6	23.1 $\pm$ 1.5	24.7 $\pm$ 0.3	23.6 $\pm$ 0.3
HABER	75.2 $\pm$ 0.9	26.5 $\pm$ 1.8	110.5 $\pm$ 1.1	26.1 $\pm$ 2.1	84.7 $\pm$ 0.8	26.1 $\pm$ 0.7	100.0 $\pm$ 0.0	27.1 $\pm$ 1.8	112.0 $\pm$ 1.0	28.4 $\pm$ 1.8	37.8 $\pm$ 0.5	26.5 $\pm$ 1.8
MUSH	12.2 $\pm$ 0.3	4.8 $\pm$ 0.5	20.6 $\pm$ 1.2	0.5 $\pm$ 0.3	12.6 $\pm$ 0.4	1.1 $\pm$ 0.1	22.0 $\pm$ 0.5	4.8 $\pm$ 0.5	59.1 $\pm$ 0.1	4.8 $\pm$ 0.5	7.9 $\pm$ 2.4	4.8 $\pm$ 0.5
PHIS	25.8 $\pm$ 0.3	11.1 $\pm$ 0.6	33.6 $\pm$ 0.2	6.4 $\pm$ 0.5	27.5 $\pm$ 0.2	6.8 $\pm$ 0.5	37.1 $\pm$ 0.5	10.4 $\pm$ 1.3	85.2 $\pm$ 0.2	11.1 $\pm$ 0.6	25.8 $\pm$ 0.4	11.1 $\pm$ 0.6
SVMG	17.4 $\pm$ 0.6	7.0 $\pm$ 1.0	28.2 $\pm$ 0.7	7.0 $\pm$ 1.0	22.0 $\pm$ 0.6	7.0 $\pm$ 1.0	29.2 $\pm$ 0.7	7.0 $\pm$ 1.0	37.6 $\pm$ 0.3	7.0 $\pm$ 1.0	8.7 $\pm$ 0.3	7.0 $\pm$ 1.0
TTT	75.9 $\pm$ 1.9	30.4 $\pm$ 3.8	100.0 $\pm$ 0.7	30.2 $\pm$ 2.7	87.2 $\pm$ 1.2	29.5 $\pm$ 4.3	94.8 $\pm$ 0.7	30.4 $\pm$ 3.8	117.0 $\pm$ 1.8	32.2 $\pm$ 2.9	69.8 $\pm$ 6.5	30.4 $\pm$ 3.8
FASHION	40.7 $\pm$ 0.2	17.1 $\pm$ 0.2	58.6 $\pm$ 0.2	13.0 $\pm$ 0.3	49.9 $\pm$ 0.3	13.0 $\pm$ 0.3	N/A	N/A	40.3 $\pm$ 0.3	18.3 $\pm$ 0.3	26.6 $\pm$ 0.1	25.4 $\pm$ 0.2
MNIST	26.9 $\pm$ 0.1	8.0 $\pm$ 0.4	44.7 $\pm$ 0.1	4.4 $\pm$ 0.1	35.8 $\pm$ 0.2	4.3 $\pm$ 0.2	N/A	N/A	37.4 $\pm$ 2.6	15.8 $\pm$ 2.7	23.6 $\pm$ 0.2	22.6 $\pm$ 0.3
PEND	12.0 $\pm$ 0.2	2.5 $\pm$ 0.3	17.1 $\pm$ 0.1	1.5 $\pm$ 0.2	15.5 $\pm$ 0.2	1.5 $\pm$ 0.2	N/A	N/A	52.7 $\pm$ 0.6	5.8 $\pm$ 0.7	10.6 $\pm$ 0.2	8.3 $\pm$ 0.7
PROTEIN	84.9 $\pm$ 0.6	34.9 $\pm$ 0.3	138.3 $\pm$ 0.3	36.6 $\pm$ 0.5	120.1 $\pm$ 1.3	55.0 $\pm$ 0.5	N/A	N/A	73.4 $\pm$ 0.3	39.7 $\pm$ 0.5	54.1 $\pm$ 0.2	52.0 $\pm$ 0.6
SENSOR	16.8 $\pm$ 0.4	4.1 $\pm$ 0.3	21.2 $\pm$ 1.0	3.9 $\pm$ 0.4	15.3 $\pm$ 1.0	3.0 $\pm$ 0.2	N/A	N/A	32.9 $\pm$ 2.0	9.6 $\pm$ 2.0	11.4 $\pm$ 1.0	10.6 $\pm$ 0.6

Empirical results

Table – Proposed approach's bound and test error for every distribution. **Bolded** values are the best bounds, underlined values are the best error. Average and standard deviations over 5 random seeds.

Task	Categorical - Part		Dirichlet - Part		Gaussian - Part	
	Bound	Test error	Bound	Test error	Bound	Test error
ADULT	46.9 ± 0.0	22.2 ± 0.1	23.2 ± 2.4	<u>21.6 ± 2.5</u>	48.7 ± 4.1	<u>21.6 ± 4.3</u>
CODRNA	24.7 ± 0.3	23.6 ± 0.3	29.4 ± 4.8	28.4 ± 4.6	39.7 ± 1.4	<u>12.0 ± 0.1</u>
HABER	37.8 ± 0.5	<u>26.5 ± 1.8</u>	84.5 ± 21.2	29.7 ± 5.9	82.3 ± 1.0	<u>26.5 ± 0.9</u>
MUSH	7.9 ± 2.4	4.8 ± 0.5	7.9 ± 1.0	4.1 ± 1.8	15.7 ± 1.3	<u>0.6 ± 0.4</u>
PHIS	25.8 ± 0.4	11.1 ± 0.6	26.8 ± 0.3	<u>6.5 ± 0.4</u>	27.2 ± 0.2	<u>6.5 ± 0.4</u>
SVMG	8.7 ± 0.3	7.0 ± 1.0	9.6 ± 1.8	<u>5.3 ± 1.4</u>	13.0 ± 0.3	6.9 ± 1.5
TTT	69.8 ± 6.5	30.4 ± 3.8	84.2 ± 1.3	29.8 ± 4.2	75.9 ± 2.2	<u>29.3 ± 2.3</u>
FASHION	26.6 ± 0.1	25.4 ± 0.2	45.9 ± 0.2	<u>17.0 ± 0.2</u>	N/A	N/A
MNIST	23.6 ± 0.2	22.6 ± 0.3	43.7 ± 0.2	<u>8.0 ± 0.4</u>	N/A	N/A
PEND	10.6 ± 0.2	8.3 ± 0.7	12.9 ± 0.3	<u>2.4 ± 0.2</u>	N/A	N/A
PROTEIN	54.1 ± 0.2	<u>52.0 ± 0.6</u>	65.5 ± 29.2	52.8 ± 0.6	N/A	N/A
SENSOR	11.4 ± 1.0	10.6 ± 0.6	12.3 ± 0.8	<u>3.0 ± 0.3</u>	N/A	N/A

Thank you for attending :)!

References I

ZANTEDESCHI, Valentina, Paul VIALARD, Emilie MORVANT, Rémi EMONET, Amaury HABRARD, Pascal GERMAIN et Benjamin GUEJ (2021). "Learning Stochastic Majority Votes by Minimizing a PAC-Bayes Generalization Bound". In : *NeurIPS*, p. 455-467.